

Case Study: Bayesian Linear Regression and Sparse Bayesian Models

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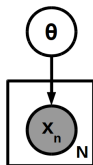
(Mini-course: lecture 2)

Nov 05, 2015

Recap

Maximum Likelihood Estimation (MLE)

- We wish to estimate parameters θ from observed data $\{\mathbf{x}_1, \dots, \mathbf{x}_N\}$



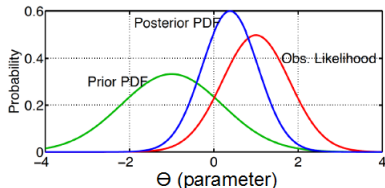
- MLE does this by finding θ that maximizes the (log)likelihood $p(\mathbf{X}|\theta)$

$$\hat{\theta} = \arg \max_{\theta} \log p(\mathbf{X}|\theta) = \arg \max_{\theta} \log \prod_{n=1}^N p(\mathbf{x}_n|\theta) = \arg \max_{\theta} \sum_{n=1}^N \log p(\mathbf{x}_n|\theta)$$

- MLE now reduces to solving an optimization problem w.r.t. θ

Maximum-a-Posteriori (MAP) Estimation

Incorporating **prior knowledge** $p(\theta)$ about the parameters



- MAP estimation finds θ that maximizes the posterior $p(\theta|\mathbf{X}) \propto p(\mathbf{X}|\theta)p(\theta)$

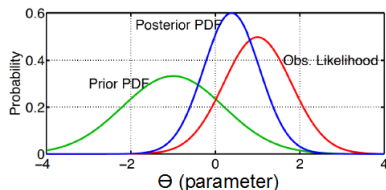
$$\hat{\theta} = \arg \max_{\theta} \log \prod_{n=1}^N p(\mathbf{x}_n|\theta)p(\theta) = \arg \max_{\theta} \sum_{n=1}^N \log p(\mathbf{x}_n|\theta) + \log p(\theta)$$

- MAP now reduces to solving an optimization problem w.r.t. θ
- Objective function very similar to MLE, except for the $\log p(\theta)$ term
- In some sense, MAP is just a “regularized” MLE

Bayesian Learning

- Both MLE and MAP only give a **point estimate** (single best answer) of θ
- How can we capture/quantify the uncertainty in θ ?
- Need to infer the **full posterior distribution**

$$p(\theta|\mathbf{X}) = \frac{p(\mathbf{X}|\theta)p(\theta)}{p(\mathbf{X})} = \frac{p(\mathbf{X}|\theta)p(\theta)}{\int_{\theta} p(\mathbf{X}|\theta)p(\theta)d\theta} \propto \text{Likelihood} \times \text{Prior}$$



- Requires doing a “fully Bayesian” inference
- Inference sometimes a somewhat easy and sometimes a (very) hard problem
- **Conjugate priors** often make life easy when doing inference

Warm-up: Least Squares Regression

- Training data: $\{\mathbf{x}_n, y_n\}_{n=1}^N$. Response is a noisy function of the input

$$y_n = f(\mathbf{x}_n, \mathbf{w}) + \epsilon_n$$

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$$E(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^N |f(\mathbf{x}_n, \mathbf{w}) - y_n|^2$$

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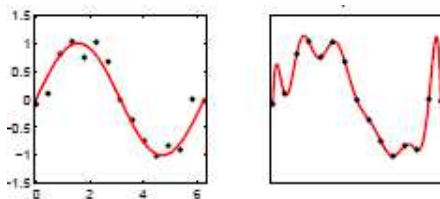
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- Classification: replace the least squares by some other loss (e.g., logistic)

Regularization

- Want functions that are “simple” (and hence “generalize” to future data)



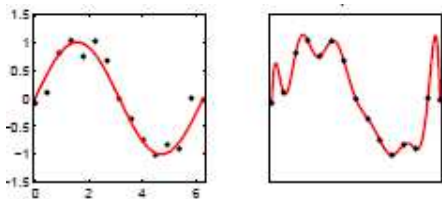
- How: penalize “complex” functions. Use a **regularized** loss function

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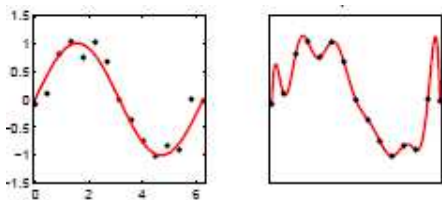
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- Regularization parameter** λ trades off **data fit** vs **model simplicity**
- For $\Omega(\mathbf{w}) = \|\mathbf{w}\|^2$, the solution $\hat{\mathbf{w}} = \arg \min_{\mathbf{w}} \tilde{E}(\mathbf{w}) = (\Phi^T \Phi + \lambda \mathbf{I})^{-1} \Phi^T \mathbf{y}$

A Probabilistic Framework for Regression

- Recall: $y_n = f(\mathbf{x}_n, \mathbf{w}) + \epsilon_n$
- Assume a zero-mean Gaussian error

$$p(\epsilon|\sigma^2) = \mathcal{N}(\epsilon|0, \sigma^2)$$

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- Leads to a Gaussian likelihood model $p(y_n|\mathbf{x}_n, \mathbf{w}) = \mathcal{N}(y_n|f(\mathbf{x}_n, \mathbf{w}), \sigma^2)$

$$p(y_n|\mathbf{x}_n, \mathbf{w}) = \left(\frac{1}{2\pi\sigma^2} \right)^{1/2} \exp \left\{ -\frac{1}{2\sigma^2} (f(\mathbf{x}_n, \mathbf{w}) - y_n)^2 \right\}$$

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- Joint probability of the data (likelihood)

$$L(\mathbf{w}) = \prod_{n=1}^N p(y_n|\mathbf{x}_n, \mathbf{w}) = \left(\frac{1}{2\pi\sigma^2} \right)^{N/2} \exp \left\{ -\frac{1}{2\sigma^2} \sum_{n=1}^N (f(\mathbf{x}_n, \mathbf{w}) - y_n)^2 \right\}$$

A Probabilistic Framework for Regression

- Let's look at the negative log-likelihood

$$-\log L(\mathbf{w}) = \frac{N}{2} \log \sigma^2 + \frac{N}{2} \log 2\pi + \frac{1}{2\sigma^2} \sum_{n=1}^N (f(\mathbf{x}_n, \mathbf{w}) - y_n)^2$$

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$$\hat{\mathbf{w}} = (\Phi^T \Phi)^{-1} \Phi^T \mathbf{y}$$

- Also get an estimate of error variance

$$\frac{1}{\hat{\sigma}^2} = \frac{1}{N} \sum_{n=1}^N (f(\mathbf{x}_n, \hat{\mathbf{w}}) - y_n)^2$$

Specifying a Prior and Computing the Posterior

- Let's assume a Gaussian prior on the weight vector $\mathbf{w} = [w_1, \dots, w_M]$

$$p(\mathbf{w}|\alpha) = \prod_{m=1}^M p(w_m|\alpha) = \prod_{m=1}^M \left(\frac{\alpha}{2\pi}\right)^{1/2} \exp\left(-\frac{\alpha}{2} w_m^2\right)$$

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- The posterior

$$p(\mathbf{w}|\mathbf{y}, \alpha, \sigma^2) = \frac{\text{likelihood} \times \text{prior}}{\text{normalizing factor}} = \frac{p(\mathbf{y}|\mathbf{w}, \sigma^2) \times p(\mathbf{w}|\alpha)}{p(\mathbf{y}|\alpha, \sigma^2)}$$

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- The posterior $p(\mathbf{w}|\mathbf{y}, \alpha, \sigma^2)$ will be Gaussian $\mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$

$$\begin{aligned}\boldsymbol{\mu} &= (\boldsymbol{\Phi}^\top \boldsymbol{\Phi} + \sigma^2 \alpha \mathbf{I})^{-1} \boldsymbol{\Phi}^\top \mathbf{y} \\ \boldsymbol{\Sigma} &= \sigma^2 (\boldsymbol{\Phi}^\top \boldsymbol{\Phi} + \sigma^2 \alpha \mathbf{I})^{-1}\end{aligned}$$

- Instead of a single estimate, we now have a **distribution** over \mathbf{w}

Maximizing the Posterior

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$$E_{MAP}(\mathbf{w}) = \frac{1}{2\sigma^2} \sum_{n=1}^N \{f(\mathbf{x}_n, \mathbf{w}) - y_n\}^2 + \frac{\alpha}{2} \sum_{m=1}^M w_m^2$$

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- Will lead to an identical solution as ridge-regression with $\lambda = \sigma^2\alpha$

Evolution of the Posterior

- Posterior updates have a naturally online flavor..

$$p(\mathbf{w}|y_1, y_2, y_3) \propto p(y_1, y_2, y_3|\mathbf{w})p(\mathbf{w})$$

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Let's Compare Predictions

- Ridge regression

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- True Bayesian

$$\text{prediction} = p(y_* | \mathbf{x}_*, \mathbf{y}, \mathbf{X}, \sigma^2, \alpha) = \int p(y_* | \mathbf{w}, \mathbf{x}_*, \sigma^2) p(\mathbf{w} | \mathbf{y}, \mathbf{X}, \alpha, \sigma^2) d\mathbf{w}$$

- The true Bayesian way **integrates out** or **marginalizes/averages over** the **uncertain variables** (\mathbf{w} in this case) to get a **predictive distribution**

Not Quite Done Yet..

- We haven't really averaged over all unknowns (which also include α , σ^2)
- Ideally, would like to get the posterior over all the unknowns

$$p(\mathbf{w}, \alpha, \sigma^2 | \mathbf{y}) = \frac{p(\mathbf{y} | \mathbf{w}, \sigma^2) p(\mathbf{w} | \alpha) p(\alpha) p(\sigma^2)}{p(\mathbf{y})}$$

where $p(\mathbf{y}) = \int p(\mathbf{y} | \mathbf{w}, \sigma^2) p(\mathbf{w} | \alpha) p(\alpha) p(\sigma^2) d\mathbf{w} d\alpha d\sigma^2$ (hard to compute)

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- Making prediction for new data points. The **predictive distribution**:

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- **Approx. Bayesian inference** (Type-II maximum likelihood, Laplace approximation, MCMC, variational Bayes, etc.) saves the day..

Approximating the Predictive Distribution

- Making prediction for new data points

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- Making prediction for new data points

$$\begin{aligned} p(y_*|\mathbf{y}) &= \int p(y_*|\mathbf{w}, \sigma^2) p(\mathbf{w}, \alpha, \sigma^2|\mathbf{y}) d\mathbf{w} d\alpha d\sigma^2 \\ &= \int p(y_*|\mathbf{w}, \sigma^2) p(\mathbf{w}|\alpha, \sigma^2, \mathbf{y}) p(\alpha, \sigma^2|\mathbf{y}) d\mathbf{w} d\alpha d\sigma^2 \\ &\approx \int p(y_*|\mathbf{w}, \sigma^2) p(\mathbf{w}|\alpha_{MP}, \sigma_{MP}^2, \mathbf{y}) \delta(\alpha_{MP}, \sigma_{MP}^2) d\mathbf{w} d\alpha d\sigma^2 \\ &= \int p(y_*|\mathbf{w}, \sigma^2) p(\mathbf{w}|\alpha_{MP}, \sigma_{MP}^2, \mathbf{y}) d\mathbf{w} \end{aligned}$$

- Recall: $p(\mathbf{w}|\alpha_{MP}, \sigma_{MP}^2, \mathbf{y})$ is a Gaussian; so is $p(y_*|\mathbf{w}, \sigma^2)$

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- Recall: $p(\mathbf{w}|\alpha_{MP}, \sigma_{MP}^2, \mathbf{y})$ is a Gaussian; so is $p(y_*|\mathbf{w}, \sigma^2)$
- Can thus now compute $p(y_*|\mathbf{y}) = \int p(y_*|\mathbf{w}, \sigma^2) p(\mathbf{w}|\alpha_{MP}, \sigma_{MP}^2, \mathbf{y}) d\mathbf{w}$, which is again a Gaussian $\mathcal{N}(y_*|\mu_*, \sigma_*^2)$

$$\begin{aligned} \mu_* &= f(\mathbf{x}_*, \mathbf{w}) \\ \sigma_*^2 &= \sigma_{MP}^2 + \phi(\mathbf{x}_*)^\top \Sigma \phi(\mathbf{x}_*) \end{aligned}$$

Marginal Likelihood

- Hyperparameters α, σ^2 are estimated by maximizing the **marginal likelihood**
- Marginal likelihood (averaged over the prior on \mathbf{w}) is

$$\begin{aligned} p(\mathbf{y}|\alpha, \sigma^2) &= \int p(\mathbf{y}|\mathbf{w}, \sigma^2)p(\mathbf{w}|\alpha)d\alpha \\ &= \frac{1}{(2\pi)^{N/2}}|\sigma^2\mathbf{I} + \Phi\mathbf{A}^{-1}\Phi^\top|^{-1/2}\exp\left(-\frac{1}{2}\mathbf{y}^\top(\sigma^2\mathbf{I} + \Phi\mathbf{A}^{-1}\Phi^\top)^{-1}\mathbf{y}\right) \end{aligned}$$

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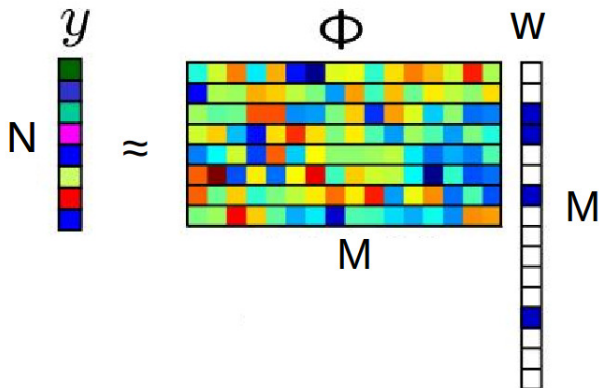
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- Maximizing $p(\mathbf{y}|\alpha, \sigma^2)$ w.r.t. α and σ^2 gives α_{MP} and σ_{MP}^2 , respectively
- Maximization can be done using gradient-based methods
- Can assume uniform priors on α, σ^2 and compute **marginal model probability**

$$\begin{aligned} p(\mathbf{y}|\mathcal{M}) &= \int p(\mathbf{y}|\alpha, \sigma^2)p(\alpha)p(\sigma^2)d\alpha d\sigma^2 \\ p(\mathbf{y}|\mathcal{M}) &\approx \frac{1}{S} \sum_{s=1}^S p(\mathbf{y}|\alpha_s, \sigma_s^2) \quad (\text{useful for model-selection}) \end{aligned}$$

Sparse Modeling



- Want very few elements in w to be nonzero

Sparse Bayesian Regression

- Recall the Gaussian prior on \mathbf{w}

$$p(\mathbf{w}|\alpha) = \prod_{m=1}^M p(w_m|\alpha) = \prod_{m=1}^M \left(\frac{\alpha}{2\pi}\right)^{1/2} \exp\left(-\frac{\alpha}{2}w_m^2\right)$$

- Each component of \mathbf{w} is a **zero-mean Gaussian** $p(w_m|\alpha) = \mathcal{N}(w_m|0, \alpha^{-1})$
- Same hyperparameter α on each entry of \mathbf{w} . Can't impose sparsity on \mathbf{w}

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- Same hyperparameter α on each entry of \mathbf{w} . Can't impose sparsity on \mathbf{w}
- Let's have a separate inverse variance α_m for each component of \mathbf{w}

$$p(\mathbf{w}|\alpha) = \prod_{m=1}^M p(w_m|\alpha_m) = \prod_{m=1}^M \left(\frac{\alpha_m}{2\pi}\right)^{1/2} \exp\left(-\frac{\alpha_m}{2}w_m^2\right)$$

- We now have M hyperparameters $\alpha = [\alpha_1, \dots, \alpha_M]$ individually controlling the variance of each component w_m of \mathbf{w}

A Hierarchical Prior

- Our new hierarchical prior on \mathbf{w}

$$p(\mathbf{w}|\alpha) = \prod_{m=1}^M p(w_m|\alpha_m) = \prod_{m=1}^M \left(\frac{\alpha_m}{2\pi}\right)^{1/2} \exp\left(-\frac{\alpha_m}{2} w_m^2\right)$$

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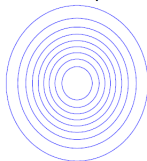
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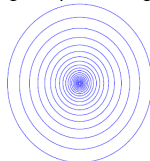
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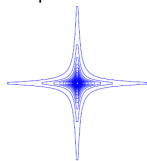
Gaussian prior



Marginal prior: single α



Independent α



A Hierarchical Prior

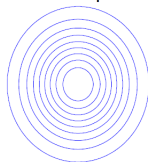
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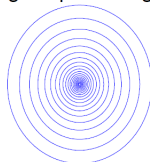
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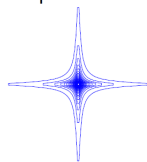
Gaussian prior



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Independent α



- Akin to penalizing $\sum_{m=1}^M \log |w_m|$. Leads to sparse solutions for \mathbf{w}

Sparse Bayesian Regression

- Likelihood model

$$p(\mathbf{y}|\mathbf{w}, \sigma^2) = (2\pi\sigma^2)^{-N/2} \exp \left\{ -\frac{1}{2\sigma^2} \|\mathbf{y} - \Phi\boldsymbol{\mu}\|^2 \right\}$$

- Prior on \mathbf{w} : Gaussian-gamma (Student-t)
- Posterior

$$p(\mathbf{w}, \alpha, \sigma^2|\mathbf{y}) = \frac{p(\mathbf{y}|\mathbf{w}, \alpha, \sigma^2)p(\mathbf{w}, \alpha, \sigma^2)}{p(\mathbf{y})}$$

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- Posterior $p(\mathbf{w}, \alpha, \sigma^2|\mathbf{y})$ is further decomposed as

$$p(\mathbf{w}, \alpha, \sigma^2|\mathbf{y}) = p(\mathbf{w}|\mathbf{y}, \alpha, \sigma^2)p(\alpha, \sigma^2|\mathbf{y})$$

The Posterior

- Posterior over weights will be Gaussian

$$\begin{aligned} p(\mathbf{w}|\mathbf{y}, \alpha, \sigma^2) &= \frac{p(\mathbf{y}|\mathbf{w}, \sigma^2)p(\mathbf{w}|\alpha)}{p(\mathbf{y}|\alpha, \sigma^2)} \\ &= (2\pi)^{(N+1)/2} |\boldsymbol{\Sigma}|^{-1/2} \exp \left\{ -\frac{1}{2} (\mathbf{w} - \boldsymbol{\mu}) \boldsymbol{\Sigma}^{-1} (\mathbf{w} - \boldsymbol{\mu}) \right\} \end{aligned}$$

where $\boldsymbol{\Sigma} = (\sigma^{-2} \boldsymbol{\Phi}^\top \boldsymbol{\Phi} + \mathbf{A})^{-1}$, $\boldsymbol{\mu} = \sigma^{-2} \boldsymbol{\Sigma} \boldsymbol{\Phi}^\top \mathbf{y}$, $\mathbf{A} = \text{diag}(\alpha_1, \alpha_2, \dots, \alpha_M)$

- Note: if $\alpha_m = \infty$ then $\mu_m = 0$

Hyperparameter Re-estimation

- Posterior over \mathbf{w} : $p(\mathbf{w}|\mathbf{y}, \alpha, \sigma^2) = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$
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- Maximize the marginal likelihood $p(\mathbf{y}|\alpha, \sigma^2)$ w.r.t. $\alpha = [\alpha_1, \dots, \alpha_M]$ and σ^2

$$\begin{aligned} \alpha_m^{new} &= \frac{\gamma_m}{\mu_m^2} \\ (\sigma^2)^{new} &= \frac{\|\mathbf{y} - \boldsymbol{\Phi}\boldsymbol{\mu}\|^2}{N - \sum_{m=1}^M \gamma_m} \end{aligned}$$

where $\gamma_m = 1 - \alpha_m \boldsymbol{\Sigma}_{mm}$

- Alternate between estimating \mathbf{w} , α , and σ^2

Approximate Bayesian Inference

Bayesian learning routinely needs to deal with intractable integrals, e.g.,

- **Normalization:** when computing the posterior distribution

$$p(\theta|\mathcal{D}) = \frac{p(\mathcal{D}|\theta)p(\theta)}{p(\mathcal{D})} = \frac{p(\mathcal{D}|\theta)p(\theta)}{\int p(\mathcal{D}|\theta)p(\theta)d\theta}$$

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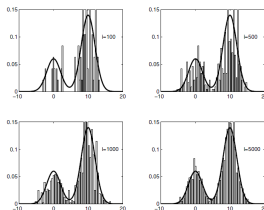
$$\mathbb{E}_{p(\theta|\mathcal{D})}[f(\mathbf{x})] = \int f(\mathbf{x})p(\theta|\mathcal{D})d\theta$$

Approximate Bayesian Inference

- Several ways to do approximate inference in Bayesian models

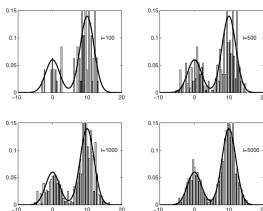
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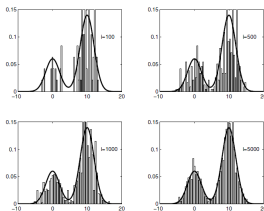
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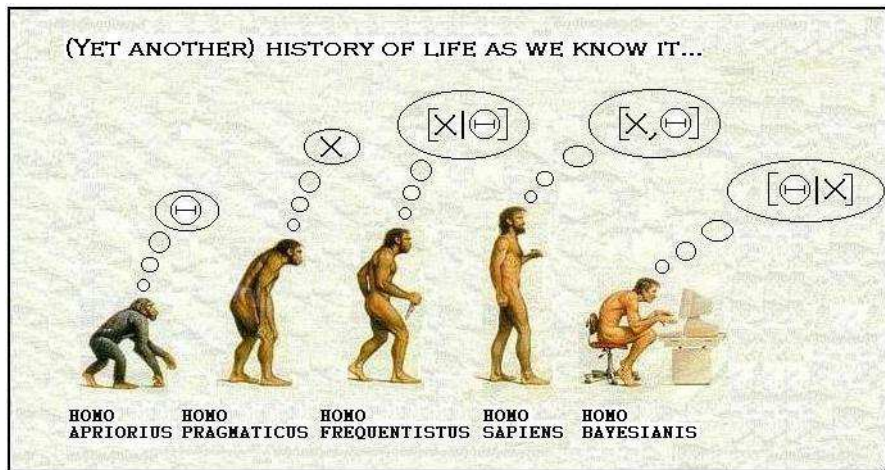
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- A very active area of research, lot of recent work on scalable inference (online and distributed Bayesian inference)

Being Bayesian



Other Recent Advances in Bayesian Learning

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- Nonparametric Bayesian modeling (or “letting the data speak for itself”)

Next Talk

- Introduction to nonparametric Bayesian modeling
- Nonparametric Bayesian regression: Gaussian Process (GP) regression

Thanks! Questions?